

BAYESIAN NETWORKS

AIMA2E CHAPTER 14.1–5 (SOME TOPICS EXCLUDED)

Outline

- ◇ Syntax
- ◇ Semantics
- ◇ Parameterized distributions
- ◇ Inference
 - Exact inference (enumeration, variable elimination)
 - Approximate inference (stochastic simulation)

Bayesian networks

A simple, graphical notation for conditional independence assertions

and hence for compact specification of full joint distributions

Syntax:

Random Variables: a set of nodes, one per variable

Topology: a directed, acyclic graph (link $\approx$ “directly influences”)

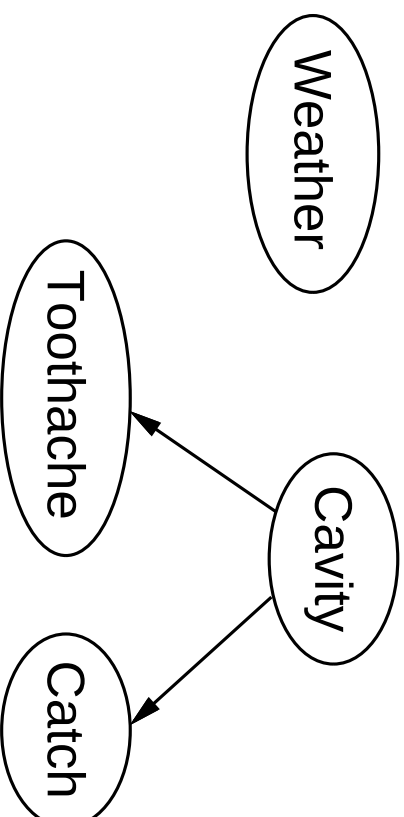
Probabilities: a conditional distribution for each node given its parents:

$$\mathbf{P}(X_i | \textit{Parents}(X_i))$$

In the simplest case, conditional distribution represented as a **conditional probability table** (CPT) giving the distribution over X_i for each combination of parent values

Example

Topology of network encodes conditional independence assertions:



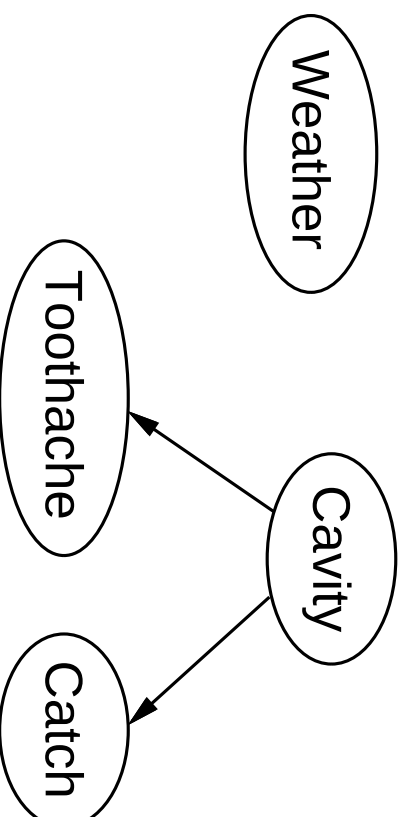
Weather is independent of the other variables

Toothache and *Catch* are conditionally independent given *Cavity*

$P(\textit{Toothache}, \textit{catch}, \textit{Cavity}, \textit{Weather}) = ?$

Example

Topology of network encodes conditional independence assertions:



Weather is independent of the other variables

Toothache and *Catch* are conditionally independent given *Cavity*

$$P(\textit{Toothache}, \textit{Catch}, \textit{Cavity}, \textit{Weather}) = P(\textit{Toothache}|\textit{Cavity})$$

$$P(\textit{Catch}|\textit{Cavity})P(\textit{Cavity})P(\textit{Weather})$$

Example

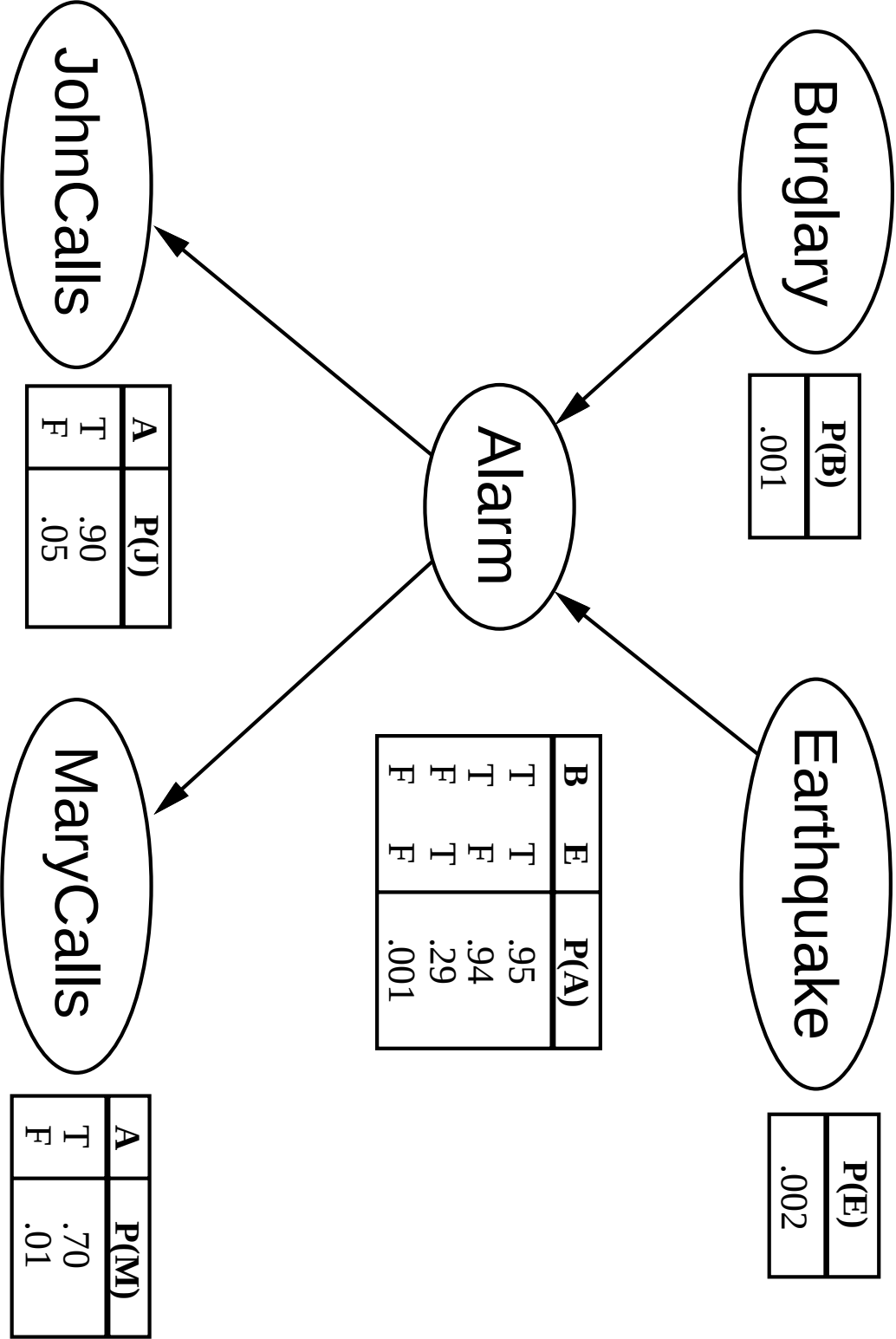
I'm at work, neighbor John calls to say my alarm is ringing, but neighbor Mary doesn't call. Sometimes it's set off by minor earthquakes. Is there a burglar?

Variables: *Burglar*, *Earthquake*, *Alarm*, *JohnCalls*, *MaryCalls*

Network topology reflects “causal” knowledge:

- A burglar can set the alarm off
- An earthquake can set the alarm off
- The alarm can cause Mary to call
- The alarm can cause John to call

Example contd.



Compactness

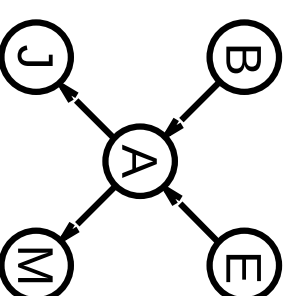
A CPT for Boolean X_i with k Boolean parents has 2^k rows for the combinations of parent values

Each row requires one number p for $X_i = true$ (the number for $X_i = false$ is just $1 - p$)

If each variable has no more than k parents, the complete network requires $O(n \cdot 2^k)$ numbers

I.e., grows linearly with n , vs. $O(2^n)$ for the full joint distribution

For burglary net, $1 + 1 + 4 + 2 + 2 = 10$ numbers (vs. $2^5 - 1 = 31$)



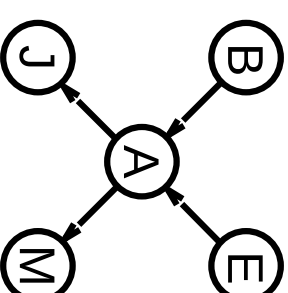
Global semantics

Global semantics defines the full joint distribution as the product of the local conditional distributions:

$$\mathbf{P}(X_1, \dots, X_n) = \prod_{i=1}^n \mathbf{P}(X_i | \text{Parents}(X_i))$$

e.g., $P(j \wedge m \wedge a \wedge \neg b \wedge \neg e)$

=



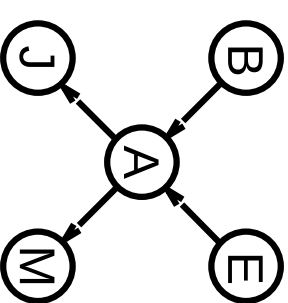
Global semantics

“Global” semantics defines the full joint distribution as the product of the local conditional distributions:

$$\mathbf{P}(X_1, \dots, X_n) = \prod_{i=1}^n \mathbf{P}(X_i | \text{Parents}(X_i))$$

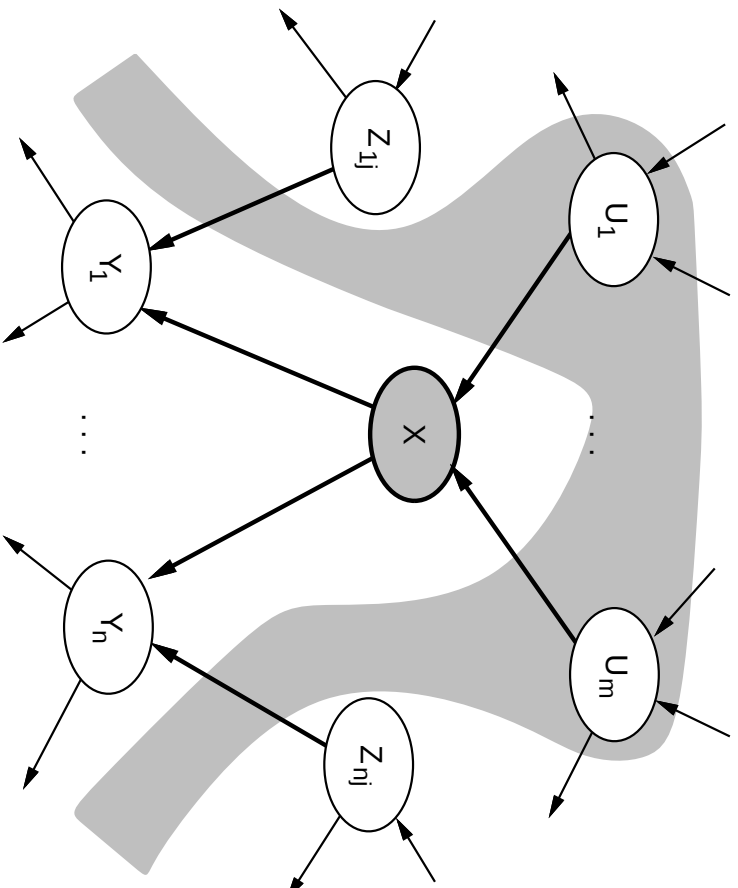
e.g., $P(j \wedge m \wedge a \wedge \neg b \wedge \neg e)$

$$= P(j|a)P(m|a)P(a|\neg b, \neg e)P(\neg b)P(\neg e)$$



Local semantics

Local semantics: each node is conditionally independent of its nondescendants given its parents

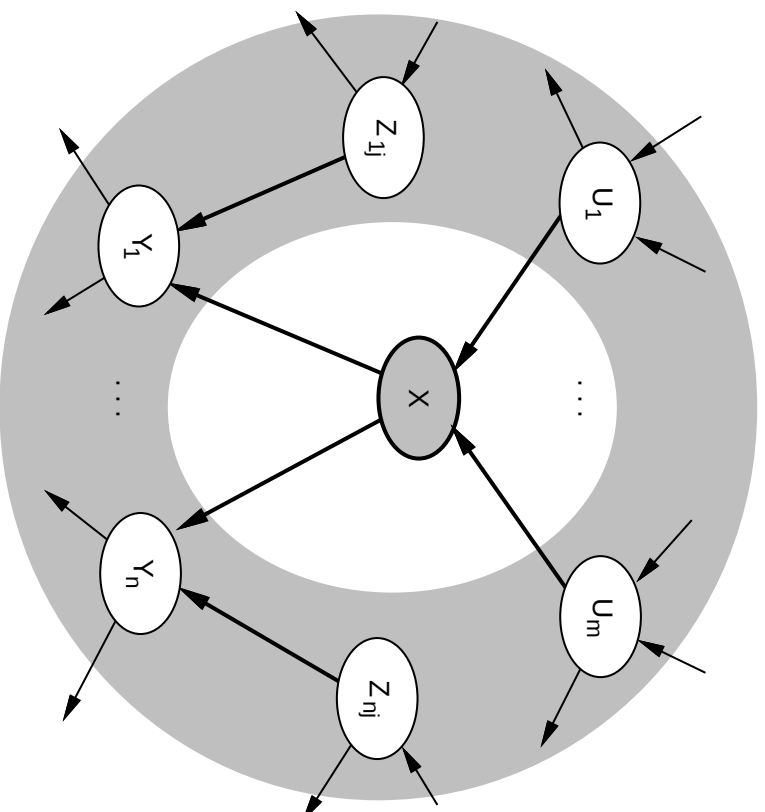


Theorem: **Local semantics** $\Leftrightarrow$ **global semantics**

Markov blanket

Each node is conditionally independent of all others given its

Markov blanket: parents + children + children's parents



Constructing Bayesian networks

Need a method such that a series of locally testable assertions of conditional independence guarantees the required global semantics

1. Choose an ordering of variables $X_1, \dots, X_n$
2. For $i = 1$ to n

add X_i to the network

select parents from $X_1, \dots, X_{i-1}$ such that

$$\mathbf{P}(X_i | Parents(X_i)) = \mathbf{P}(X_i | X_1, \dots, X_{i-1})$$

This choice of parents guarantees the global semantics:

$$\begin{aligned} \mathbf{P}(X_1, \dots, X_n) &= \prod_{i=1}^n \mathbf{P}(X_i | X_1, \dots, X_{i-1}) && \text{(chain rule)} \\ &= \prod_{i=1}^n \mathbf{P}(X_i | Parents(X_i)) && \text{(by construction)} \end{aligned}$$

Example

Suppose we choose the ordering M, J, A, B, E

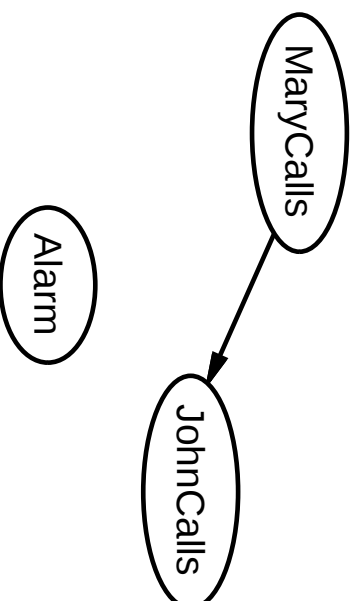
MaryCalls

JohnCalls

$$P(J|M) = P(J)?$$

Example

Suppose we choose the ordering M, J, A, B, E

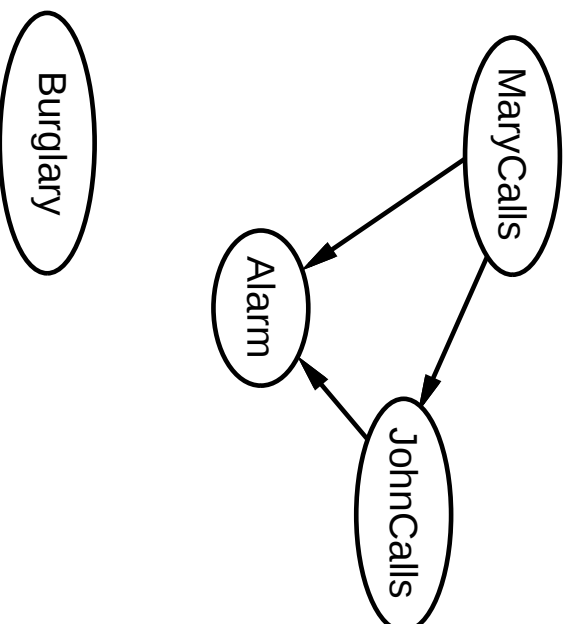


$P(J|M) = P(J)$? No

$P(A|J, M) = P(A|J)$? $P(A|J, M) = P(A)$?

Example

Suppose we choose the ordering M, J, A, B, E



$P(J|M) = P(J)$? No

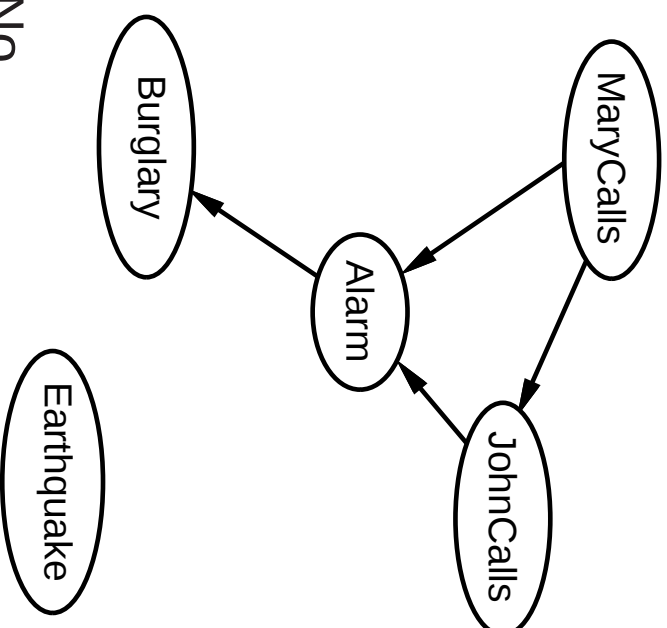
$P(A|J, M) = P(A|J)$? $P(A|J, M) = P(A)$? No

$P(B|A, J, M) = P(B|A)$?

$P(B|A, J, M) = P(B)$?

Example

Suppose we choose the ordering M, J, A, B, E



$P(J|M) = P(J)$? No

$P(A|J, M) = P(A|J)$? $P(A|J, M) = P(A)$? No

$P(B|A, J, M) = P(B|A)$? Yes

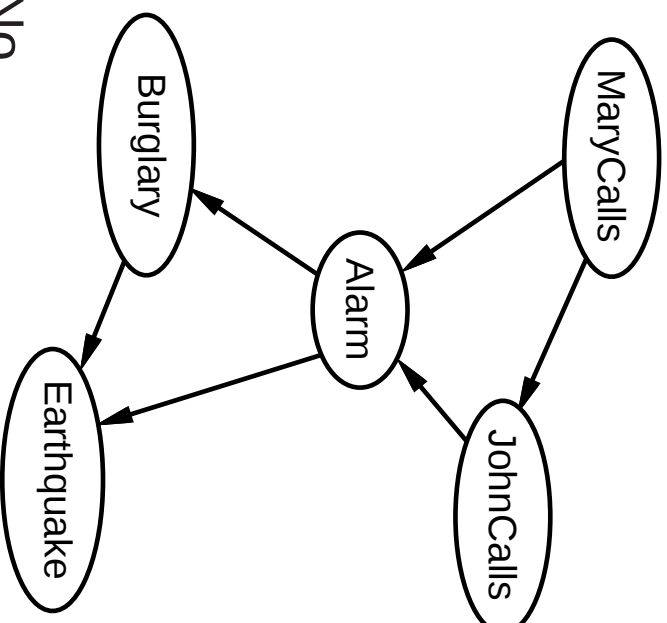
$P(B|A, J, M) = P(B)$? No

$P(E|B, A, J, M) = P(E|A)$?

$P(E|B, A, J, M) = P(E|A, B)$?

Example

Suppose we choose the ordering M, J, A, B, E



$P(J|M) = P(J)$? No

$P(A|J, M) = P(A|J)$? $P(A|J, M) = P(A)$? No

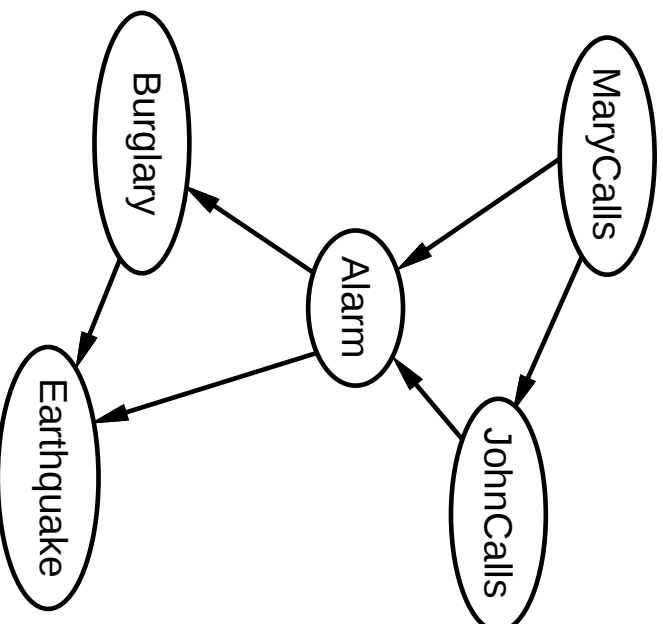
$P(B|A, J, M) = P(B|A)$? Yes

$P(B|A, J, M) = P(B)$? No

$P(E|B, A, J, M) = P(E|A)$? No

$P(E|B, A, J, M) = P(E|A, B)$? Yes

Example contd.



Deciding conditional independence is hard in noncausal directions
(Causal models and conditional independence seem hardwired for humans!)

Assessing conditional probabilities is hard in noncausal directions

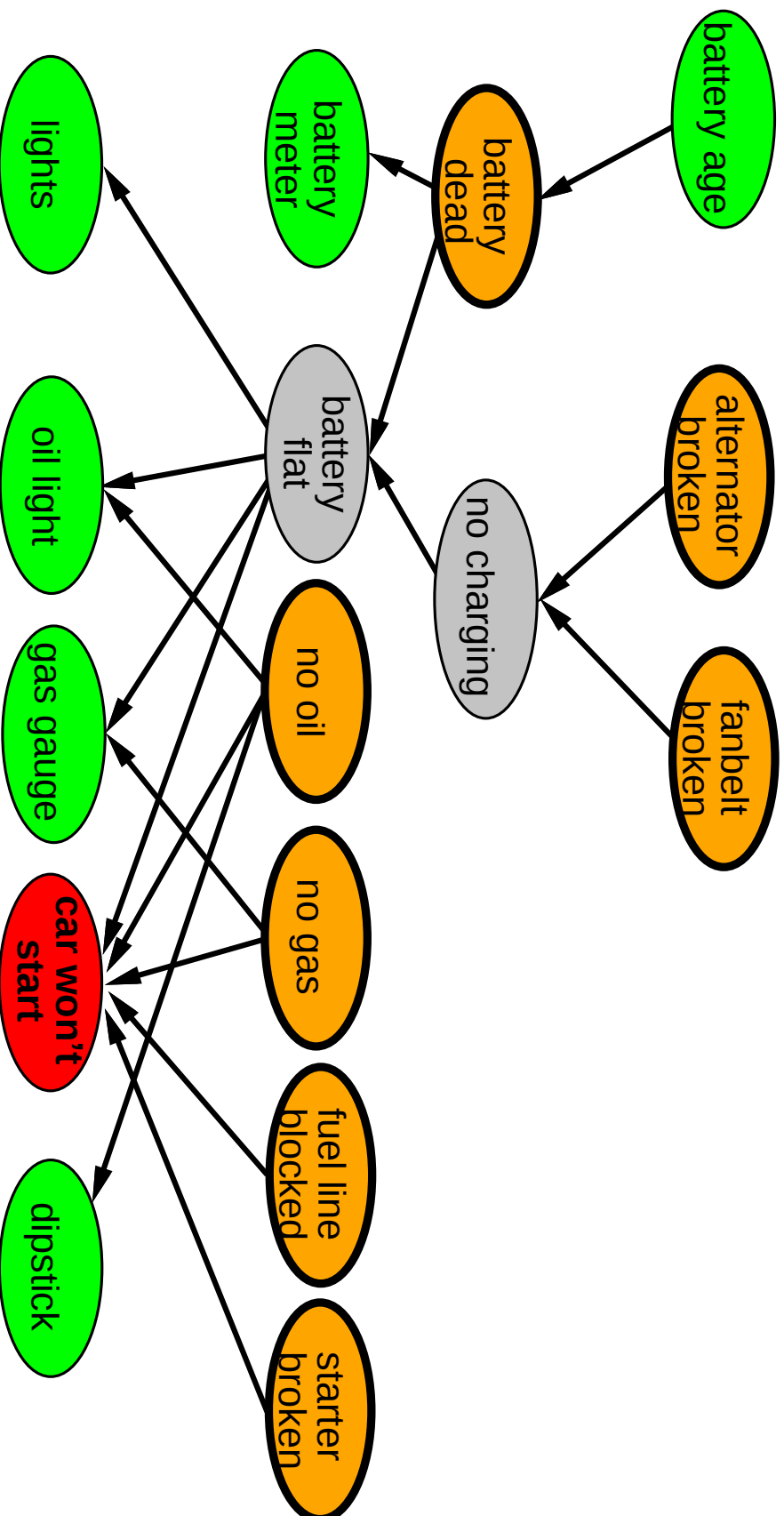
Network is less compact: $1 + 2 + 2 + 4 + 2 + 4 = 13$ numbers needed

Example: Car diagnosis

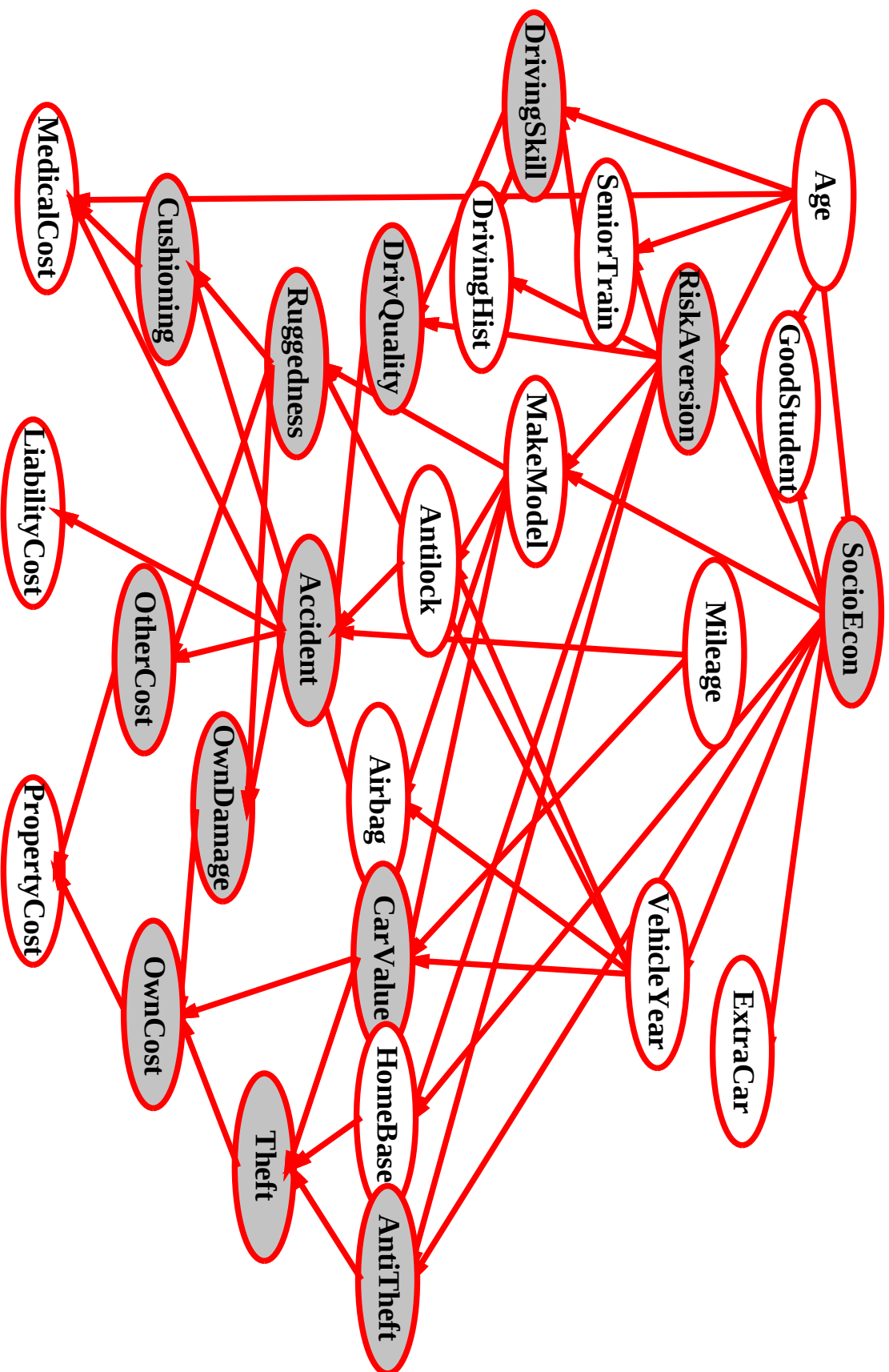
Initial evidence: car won't start

Testable variables (green), "broken, so fix it" variables (orange)

Hidden variables (gray) ensure sparse structure, reduce parameters



Example: Car insurance



Compact conditional distributions

CPT grows exponentially with no. of parents

CPT becomes infinite with continuous-valued parent or child

Solution: **canonical** distributions that are defined compactly

Deterministic nodes are the simplest case:

$$X = f(Parents(X)) \text{ for some function } f$$

E.g., Boolean functions

$$NorthAmerican \Leftrightarrow Canadian \vee US \vee Mexican$$

E.g., numerical relationships among continuous variables

$$\frac{\partial Level}{\partial t} = \text{inflow} + \text{precipitation} - \text{outflow} - \text{evaporation}$$

Compact conditional distributions contd.

Noisy-OR distributions model multiple noninteracting causes

- 1) Parents $U_1 \dots U_k$ include all causes (can add **leak node**)
- 2) Independent failure probability q_i for each cause alone

$$\Rightarrow P(X|U_1 \dots U_j, \neg U_{j+1} \dots \neg U_k) = 1 - \prod_{i=1}^j q_i$$

<i>Cold</i>	<i>Flu</i>	<i>Malaria</i>	$P(\text{Fever})$	$P(\neg \text{Fever})$
F	F	F	0.0	1.0
F	F	T	0.9	0.1
F	T	F	0.8	0.2
F	T	T	0.98	$0.02 = 0.2 \times 0.1$
T	F	F	0.4	0.6
T	F	T	0.94	$0.06 = 0.6 \times 0.1$
T	T	F	0.88	$0.12 = 0.6 \times 0.2$
T	T	T	0.988	$0.012 = 0.6 \times 0.2 \times 0.1$

Number of parameters **linear** in number of parents

Continuous nodes

Networks may have discrete RVs, continuous RVs, or a mix of the two.

All continuous (e.g., conditional Gaussian). Linear dynamic systems (Kalman filter).

Discrete parents, continuous children (e.g., conditional Gaussian). Gaussian mixture models.

Continuous parents, discrete children (e.g., probit and logit functions). Difficult to deal with.

Inference tasks

Simple queries: compute posterior marginal $\mathbf{P}(X_i|\mathbf{E}=\mathbf{e})$

e.g., $P(NoGas|Gauge = empty, Lights = on, Starts = false)$

Conjunctive queries: $\mathbf{P}(X_i, X_j|\mathbf{E}=\mathbf{e}) = \mathbf{P}(X_i|\mathbf{E}=\mathbf{e})\mathbf{P}(X_j|X_i, \mathbf{E}=\mathbf{e})$

Optimal decisions: decision networks include utility information;
probabilistic inference required for $P(outcome|action, evidence)$

Value of information: which evidence to seek next?

Sensitivity analysis: which probability values are most critical?

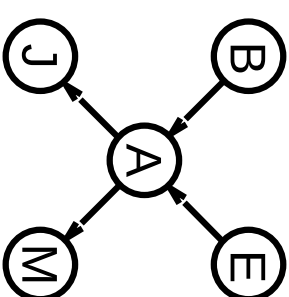
Explanation: why do I need a new starter motor?

Inference by enumeration

Slightly intelligent way to sum out variables from the joint without actually constructing its explicit representation

Simple query on the burglary network:

$$\begin{aligned}
 & \mathbf{P}(B|j, m) \\
 &= \mathbf{P}(B, j, m) / P(j, m) \\
 &= \alpha \mathbf{P}(B, j, m) \\
 &= \alpha \sum_e \sum_a \mathbf{P}(B, e, a, j, m)
 \end{aligned}$$



Rewrite full joint entries using product of CPT entries:

$$\begin{aligned}
 & \mathbf{P}(B|j, m) \\
 &= \alpha \sum_e \sum_a \mathbf{P}(B) P(e) \mathbf{P}(a|B, e) P(j|a) P(m|a) \\
 &= \alpha \mathbf{P}(B) \sum_e P(e) \sum_a \mathbf{P}(a|B, e) P(j|a) P(m|a)
 \end{aligned}$$

Recursive depth-first enumeration: $O(n)$ space, $O(d^n)$ time

Enumeration algorithm

```
function ENUMERATION-ASK( $X, \mathbf{e}, bn$ ) returns a distribution over  $X$ 
inputs:  $X$ , the query variable
 $\mathbf{e}$ , observed values for variables  $\mathbf{E}$ 
 $bn$ , a Bayesian network with variables  $\{X\} \cup \mathbf{E} \cup \mathbf{Y}$ 

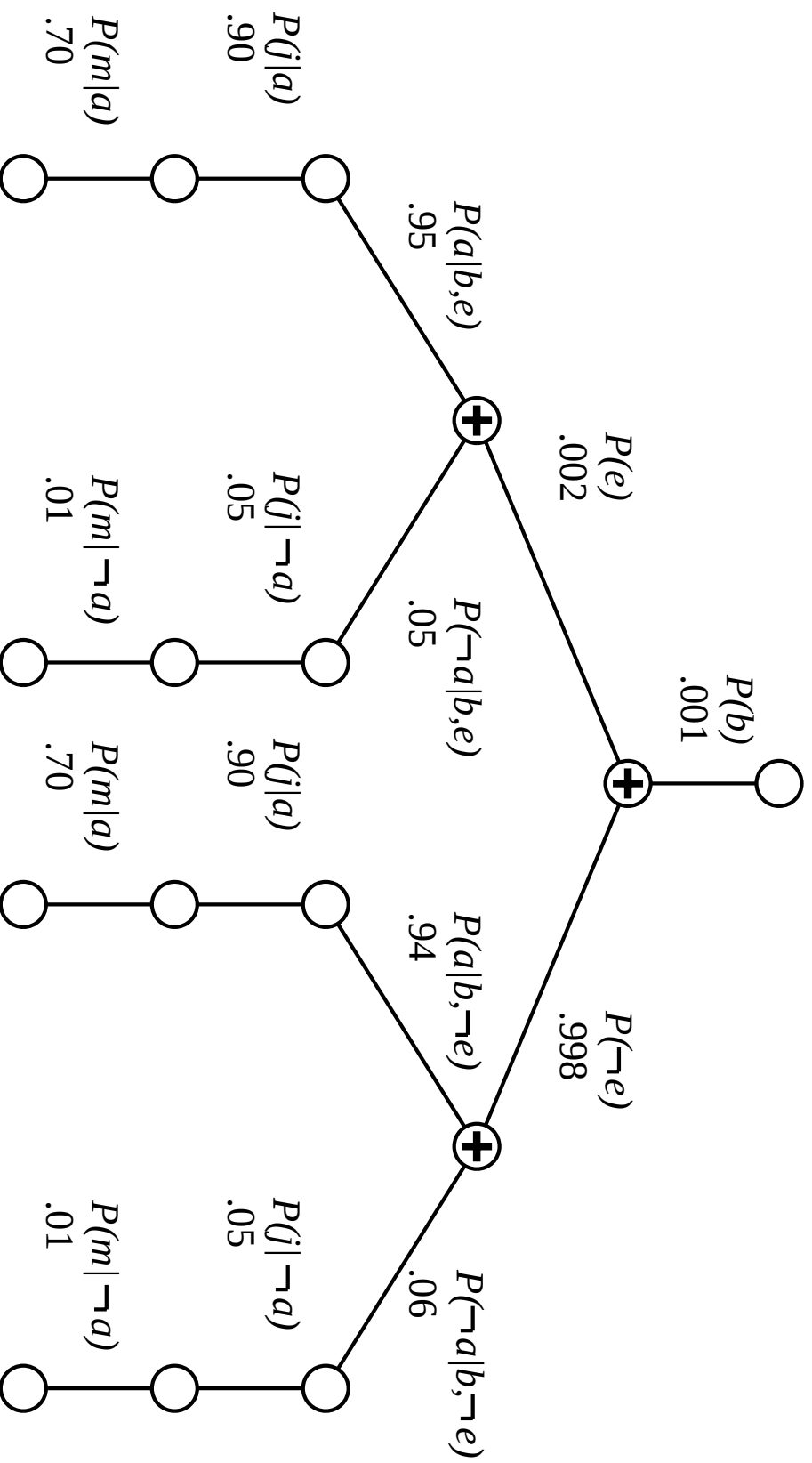
 $Q(X) \leftarrow$  a distribution over  $X$ , initially empty
for each value  $x_i$  of  $X$  do
  extend  $\mathbf{e}$  with value  $x_i$  for  $X$ 
   $Q(x_i) \leftarrow$  ENUMERATE-ALL(VARS[ $bn$ ],  $\mathbf{e}$ )
return NORMALIZE( $Q(X)$ )
```

```
function ENUMERATE-ALL( $vars, \mathbf{e}$ ) returns a real number
if EMPTY?( $vars$ ) then return 1.0
 $Y \leftarrow$  FIRST( $vars$ )
if  $Y$  has value  $y$  in  $\mathbf{e}$ 
  then return  $P(y \mid Pa(Y)) \times$  ENUMERATE-ALL(REST( $vars$ ),  $\mathbf{e}$ )
else return  $\sum_y P(y \mid Pa(Y)) \times$  ENUMERATE-ALL(REST( $vars$ ),  $\mathbf{e}_y$ )
  where  $\mathbf{e}_y$  is  $\mathbf{e}$  extended with  $Y = y$ 
```

Evaluation tree

Enumeration is inefficient: repeated computation

e.g., computes $P(j|a)P(m|a)$ for each value of e



Inference by variable elimination

Variable elimination: carry out summations right-to-left, storing intermediate results (**factors**) to avoid recomputation

$$\begin{aligned}
 & \mathbf{P}(B|j, m) \\
 &= \alpha \underbrace{\mathbf{P}(B)}_B \underbrace{\sum_e P(e)}_E \underbrace{\sum_a \mathbf{P}(a|B, e)}_A \underbrace{P(j|a)}_j \underbrace{P(m|a)}_M \\
 &= \alpha \mathbf{P}(B) \sum_e P(e) \sum_a \mathbf{P}(a|B, e) P(j|a) f_M(a) \\
 &= \alpha \mathbf{P}(B) \sum_e P(e) \sum_a \mathbf{P}(a|B, e) f_J(a) f_M(a) \\
 &= \alpha \mathbf{P}(B) \sum_e P(e) \sum_a f_A(a, b, e) f_J(a) f_M(a) \\
 &= \alpha \mathbf{P}(B) \sum_e P(e) f_{\bar{A}JM}(b, e) \text{ (sum out } A) \\
 &= \alpha \mathbf{P}(B) f_{\bar{E}\bar{A}JM}(b) \text{ (sum out } E) \\
 &= \alpha f_B(b) \times f_{\bar{E}\bar{A}JM}(b)
 \end{aligned}$$

Variable elimination: Basic operations

Summing out a variable from a product of factors:

move any constant factors outside the summation

add up submatrices in pointwise product of remaining factors

$$\sum_x f_1 \times \cdots \times f_k = f_1 \times \cdots \times f_i \sum_x f_{i+1} \times \cdots \times f_k = f_1 \times \cdots \times f_i \times f_{\bar{X}}$$

assuming $f_1, \dots, f_i$ do not depend on X

Pointwise product of factors f_1 and f_2 :

$$\begin{aligned} f_1(x_1, \dots, x_j, y_1, \dots, y_k) \times f_2(y_1, \dots, y_k, z_1, \dots, z_l) \\ = f(x_1, \dots, x_j, y_1, \dots, y_k, z_1, \dots, z_l) \end{aligned}$$

$$\text{E.g., } f_1(a, b) \times f_2(b, c) = f(a, b, c)$$

Variable elimination algorithm

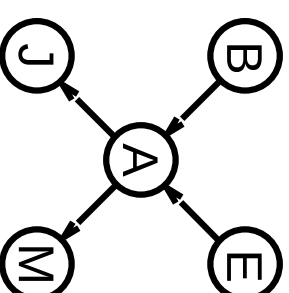
```
function ELIMINATION-ASK( $X, \mathbf{e}, bn$ ) returns a distribution over  $X$   
inputs:  $X$ , the query variable  
          $\mathbf{e}$ , evidence specified as an event  
          $bn$ , a belief network specifying joint distribution  $\mathbf{P}(X_1, \dots, X_n)$   
 $factors \leftarrow []$ ;  $vars \leftarrow \text{REVERSE}(\text{VARS}[bn])$   
for each  $var$  in  $vars$  do  
     $factors \leftarrow [\text{MAKE-FACTOR}(var, \mathbf{e}) || factors]$   
    if  $var$  is a hidden variable then  $factors \leftarrow \text{SUM-OUT}(var, factors)$   
return  $\text{NORMALIZE}(\text{POINTWISE-PRODUCT}(factors))$ 
```

Irrelevant variables

Consider the query $P(\textit{JohnCalls} | \textit{Burglary} = \textit{true})$

$$P(J|b) = \alpha P(b) \sum_e P(e) \sum_a P(a|b, e) P(J|a) \sum_m P(m|a)$$

Sum over m is identically 1; M is **irrelevant** to the query



Thm 1: Y is irrelevant unless $Y \in \textit{Ancestors}(\{X\} \cup \mathbf{E})$

Here, $X = \textit{JohnCalls}$, $\mathbf{E} = \{\textit{Burglary}\}$, and $\textit{Ancestors}(\{X\} \cup \mathbf{E}) = \{\textit{Alarm}, \textit{Earthquake}\}$ so M is irrelevant

Complexity of exact inference

Singly connected networks (or **polytrees**):

- any two nodes are connected by at most one (undirected) path
- time and space cost of variable elimination are $O(d^k n)$

Multiply connected networks:

- can reduce 3SAT to exact inference $\Rightarrow$ NP-hard
- equivalent to **counting** 3SAT models $\Rightarrow$ #P-complete

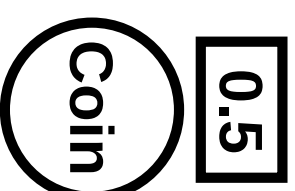
Inference by stochastic simulation

Basic idea:

- 1) Draw N samples from a sampling distribution S
- 2) Compute an approximate posterior probability $\hat{P}$
- 3) Show this converges to the true probability P

Outline:

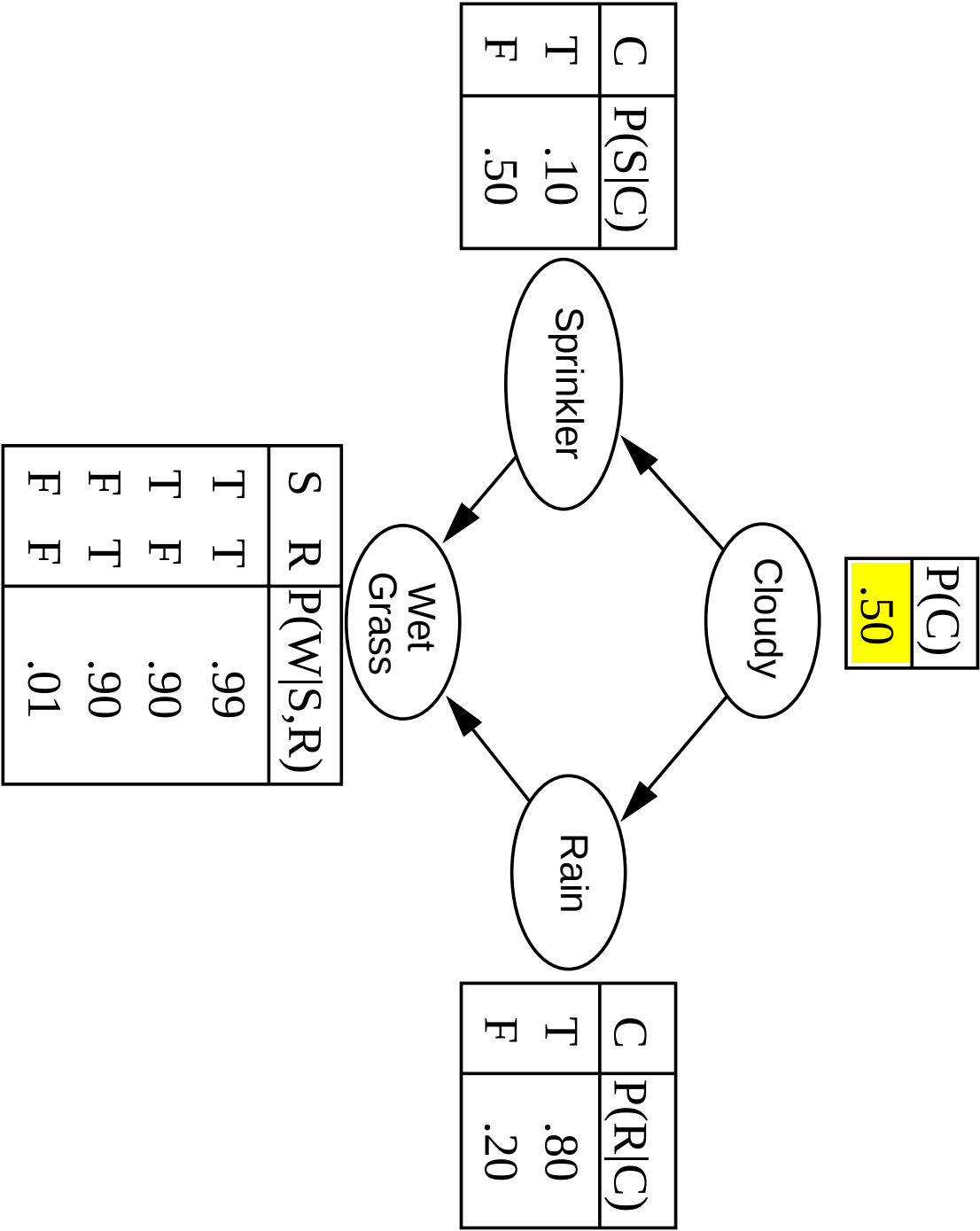
- Sampling from an empty network
- Rejection sampling: reject samples disagreeing with evidence
- Likelihood weighting: use evidence to weight samples



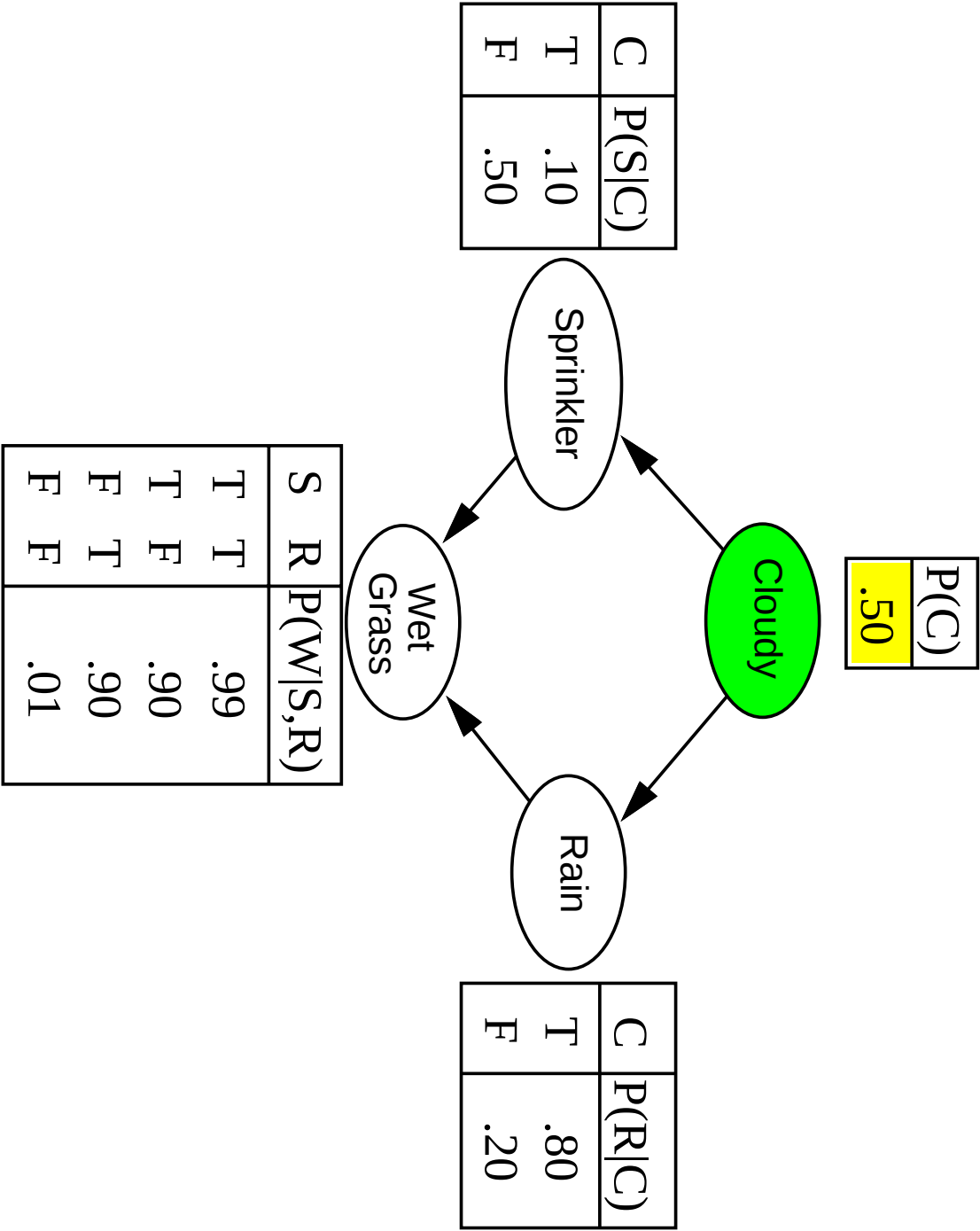
Sampling from an empty network

```
function PRIOR-SAMPLE( $bn$ ) returns an event sampled from  $bn$   
inputs:  $bn$ , a belief network specifying joint distribution  $\mathbf{P}(X_1, \dots, X_n)$   
 $\mathbf{x} \leftarrow$  an event with  $n$  elements  
for  $i = 1$  to  $n$  do  
     $x_i \leftarrow$  a random sample from  $\mathbf{P}(X_i \mid Parents(X_i))$   
return  $\mathbf{x}$ 
```

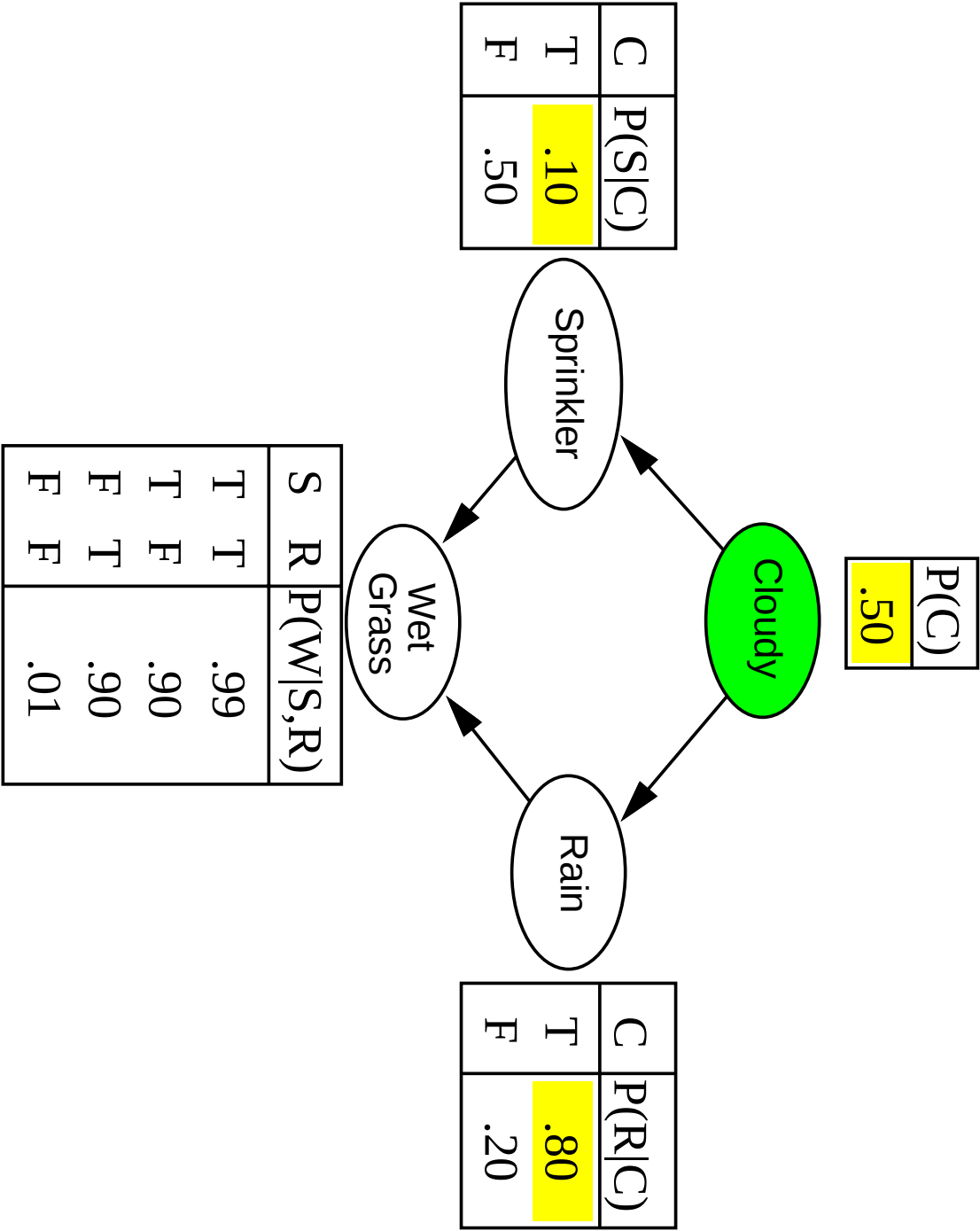
Example



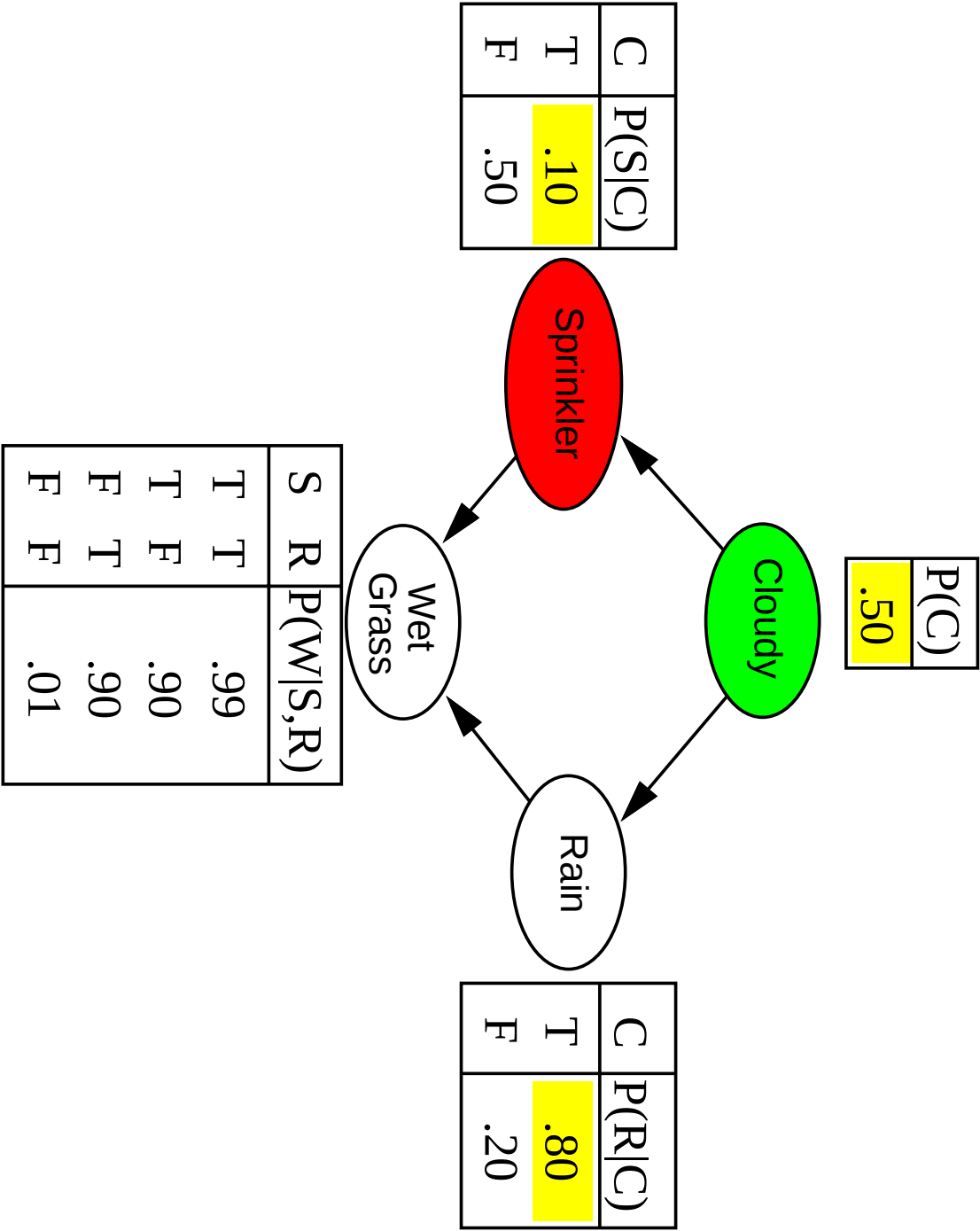
Example



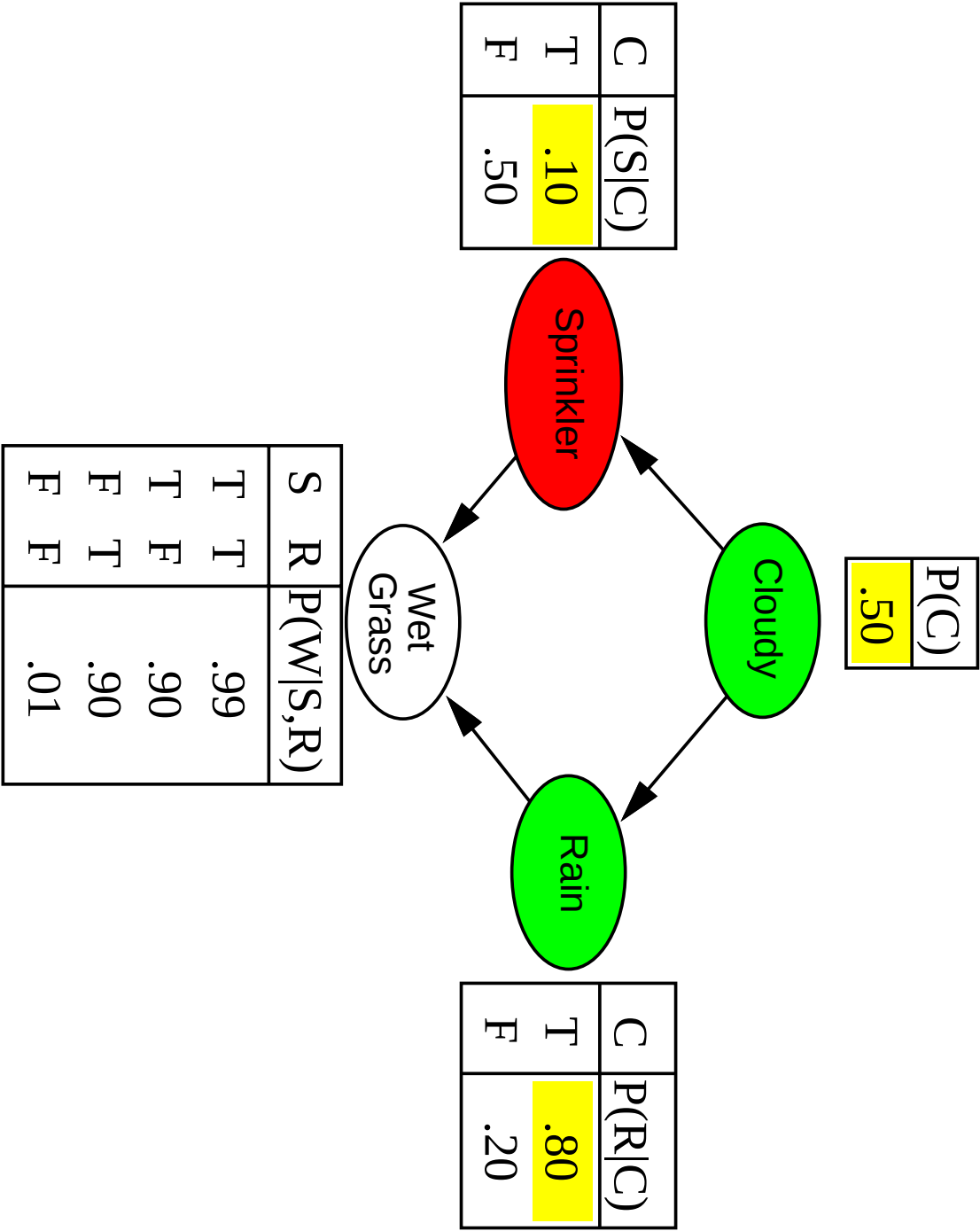
Example



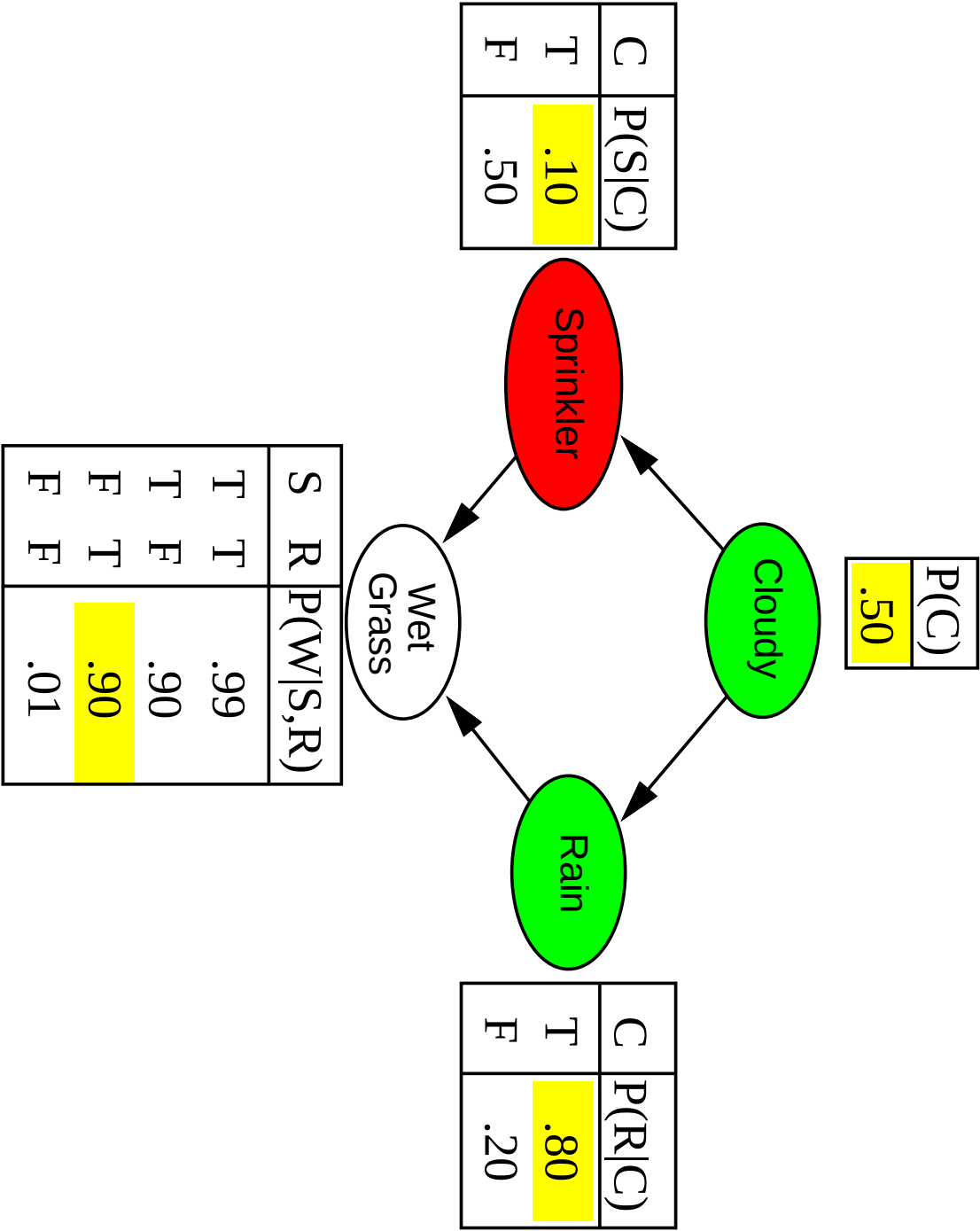
Example



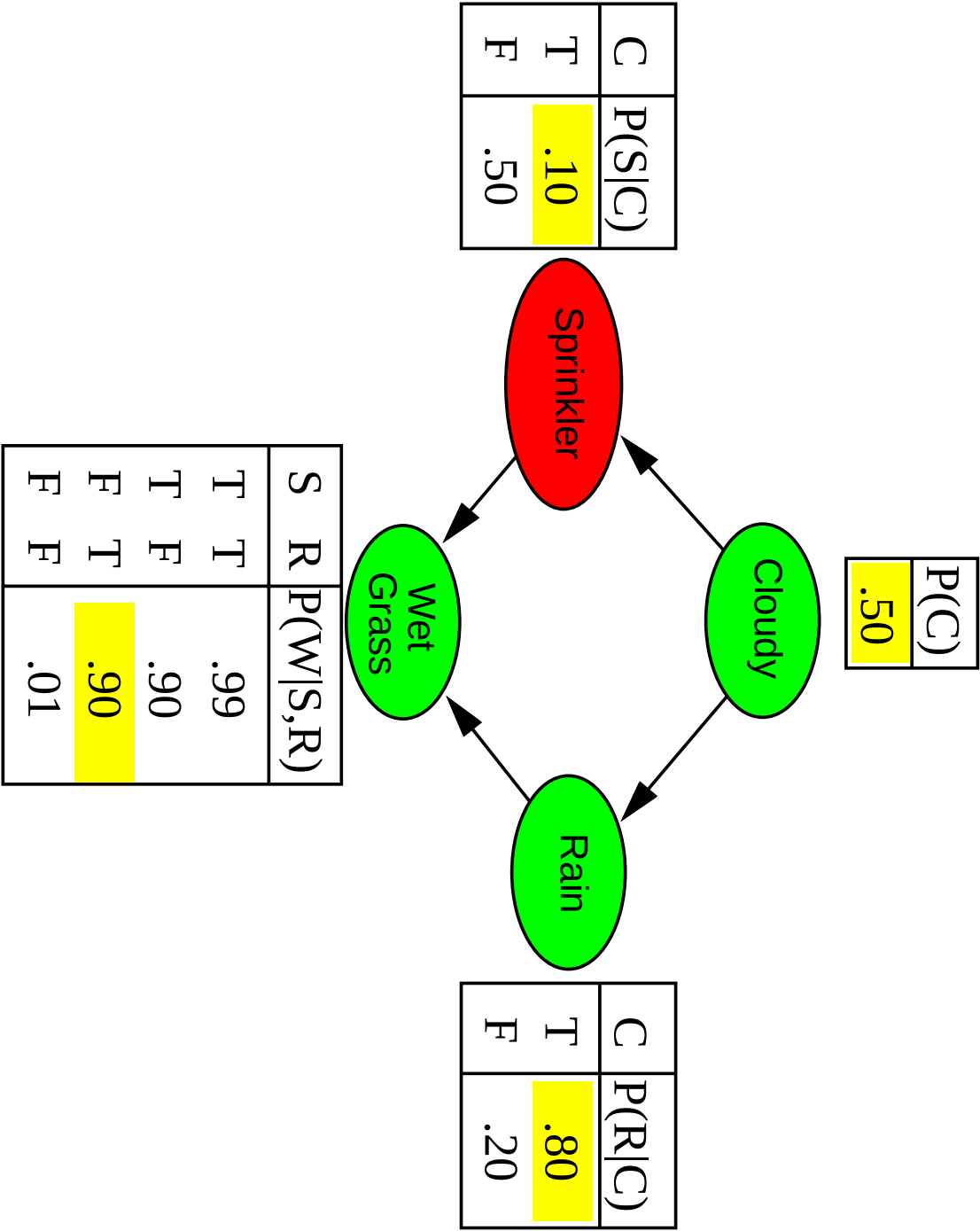
Example



Example



Example



Sampling from an empty network contd.

Probability that PRIORSAMPLE generates a particular event

$$S_{PS}(x_1 \dots x_n) = \prod_{i=1}^n P(x_i | Parents(X_i)) = P(x_1 \dots x_n)$$

i.e., the true prior probability

E.g., $S_{PS}(t, f, t, t) = 0.5 \times 0.9 \times 0.8 \times 0.9 = 0.324 = P(t, f, t, t)$

Let $N_{PS}(x_1 \dots x_n)$ be the number of samples generated for event $x_1, \dots, x_n$

Then we have

$$\begin{aligned} \lim_{N \rightarrow \infty} \hat{P}(x_1, \dots, x_n) &= \lim_{N \rightarrow \infty} N_{PS}(x_1, \dots, x_n) / N \\ &= S_{PS}(x_1, \dots, x_n) \\ &= P(x_1 \dots x_n) \end{aligned}$$

That is, estimates derived from PRIORSAMPLE are **consistent**

Shorthand: $\hat{P}(x_1, \dots, x_n) \approx P(x_1 \dots x_n)$

Rejection sampling

$\hat{\mathbf{P}}(X|\mathbf{e})$ estimated from samples agreeing with $\mathbf{e}$

```
function REJECTION-SAMPLING( $X, \mathbf{e}, bn, N$ ) returns an estimate of  $P(X|\mathbf{e})$ 
  local variables:  $\mathbf{N}$ , a vector of counts over  $X$ , initially zero
  for  $j = 1$  to  $N$  do
     $\mathbf{x} \leftarrow \text{PRIOR-SAMPLE}(bn)$ 
    if  $\mathbf{x}$  is consistent with  $\mathbf{e}$  then
       $\mathbf{N}[x] \leftarrow \mathbf{N}[x] + 1$  where  $x$  is the value of  $X$  in  $\mathbf{x}$ 
  return NORMALIZE( $\mathbf{N}[X]$ )
```

E.g., estimate $\mathbf{P}(Rain|Sprinkler = true)$ using 100 samples
27 samples have *Sprinkler = true*

Of these, 8 have *Rain = true* and 19 have *Rain = false*.

$\hat{\mathbf{P}}(Rain|Sprinkler = true) = \text{NORMALIZE}(\langle 8, 19 \rangle) = \langle 0.296, 0.704 \rangle$

Similar to a basic real-world empirical estimation procedure

Analysis of rejection sampling

$$\begin{aligned}\hat{\mathbf{P}}(X|\mathbf{e}) &= \alpha \mathbf{N}_{PS}(X, \mathbf{e}) && \text{(algorithm defn.)} \\ &= \mathbf{N}_{PS}(X, \mathbf{e}) / N_{PS}(\mathbf{e}) && \text{(normalized by } N_{PS}(\mathbf{e}) \text{)} \\ &\approx \mathbf{P}(X, \mathbf{e}) / P(\mathbf{e}) && \text{(property of PRIORSAMPLE)} \\ &= \mathbf{P}(X|\mathbf{e}) && \text{(defn. of conditional probability)}\end{aligned}$$

Hence rejection sampling returns consistent posterior estimates

Problem: hopelessly expensive if $P(\mathbf{e})$ is small

$P(\mathbf{e})$ drops off exponentially with number of evidence variables!

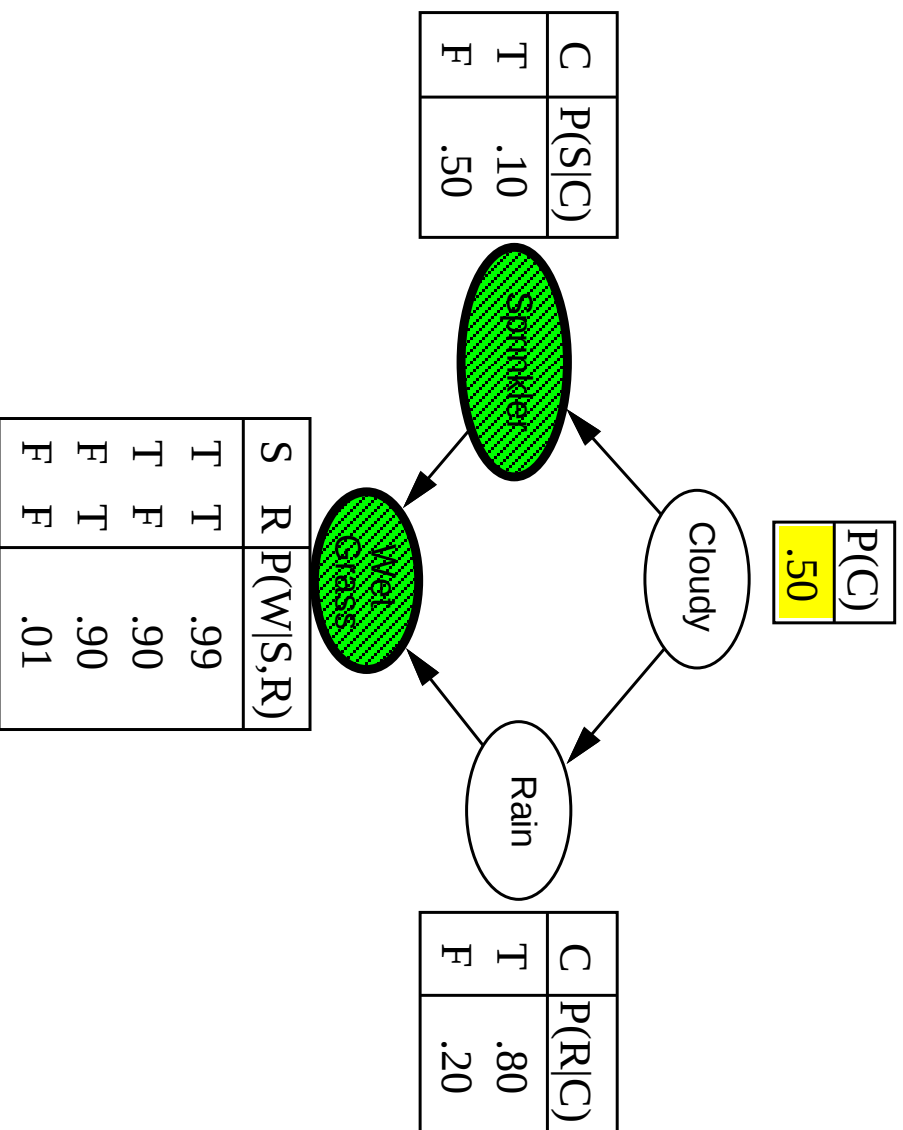
Likelihood weighting

Idea: fix evidence variables, sample only nonevidence variables, and weight each sample by the likelihood it accords the evidence

```
function LIKELIHOOD-WEIGHTING( $X, \mathbf{e}, bn, N$ ) returns an estimate of  $P(X|\mathbf{e})$   
local variables:  $\mathbf{W}$ , a vector of weighted counts over  $X$ , initially zero  
for  $j = 1$  to  $N$  do  
   $\mathbf{x}, w \leftarrow \text{WEIGHTED-SAMPLE}(bn)$   
   $\mathbf{W}[x] \leftarrow \mathbf{W}[x] + w$  where  $x$  is the value of  $X$  in  $\mathbf{x}$   
  return NORMALIZE( $\mathbf{W}[X]$ )
```

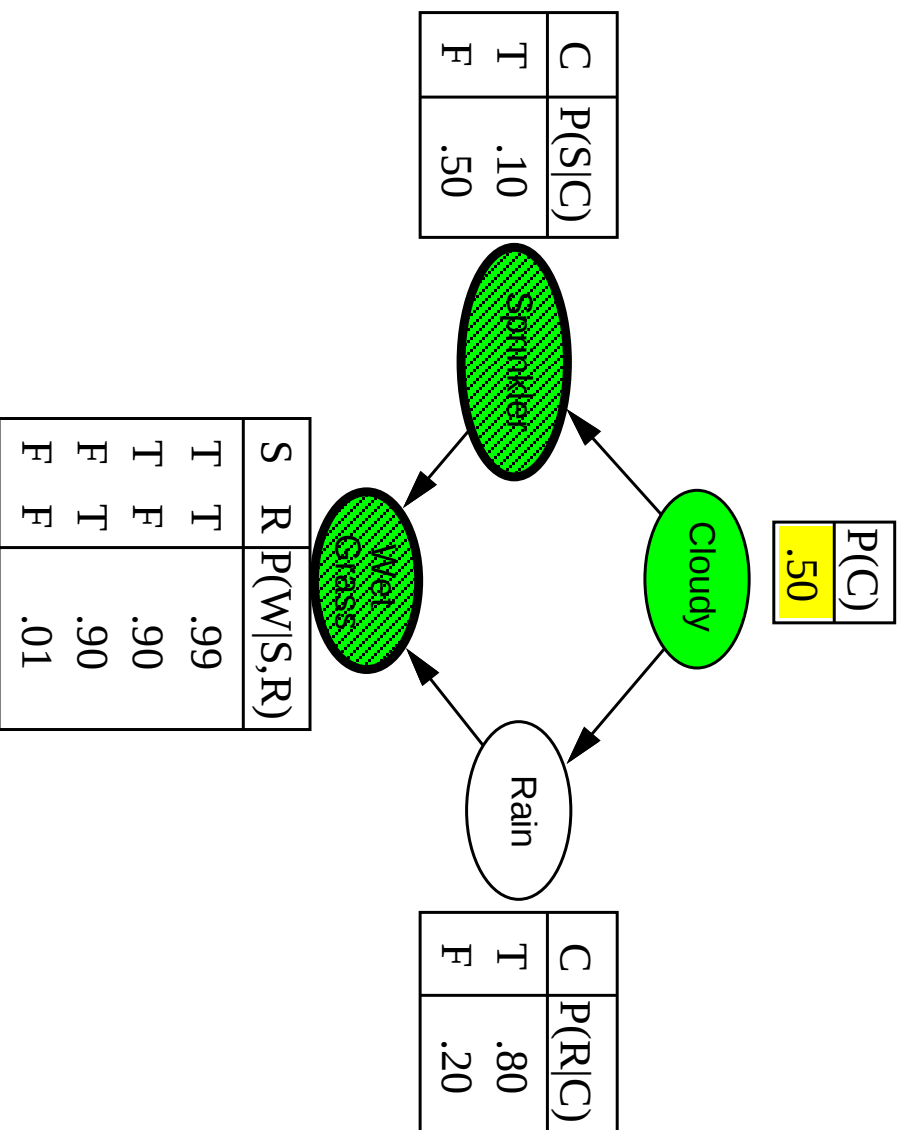
```
function WEIGHTED-SAMPLE( $bn, \mathbf{e}$ ) returns an event and a weight  
 $\mathbf{x} \leftarrow$  an event with  $n$  elements;  $w \leftarrow 1$   
for  $i = 1$  to  $n$  do  
  if  $X_i$  has a value  $x_i$  in  $\mathbf{e}$   
    then  $w \leftarrow w \times P(X_i = x_i \mid \text{Parents}(X_i))$   
    else  $x_i \leftarrow$  a random sample from  $\mathbf{P}(X_i \mid \text{Parents}(X_i))$   
return  $\mathbf{x}, w$ 
```

Likelihood weighting example



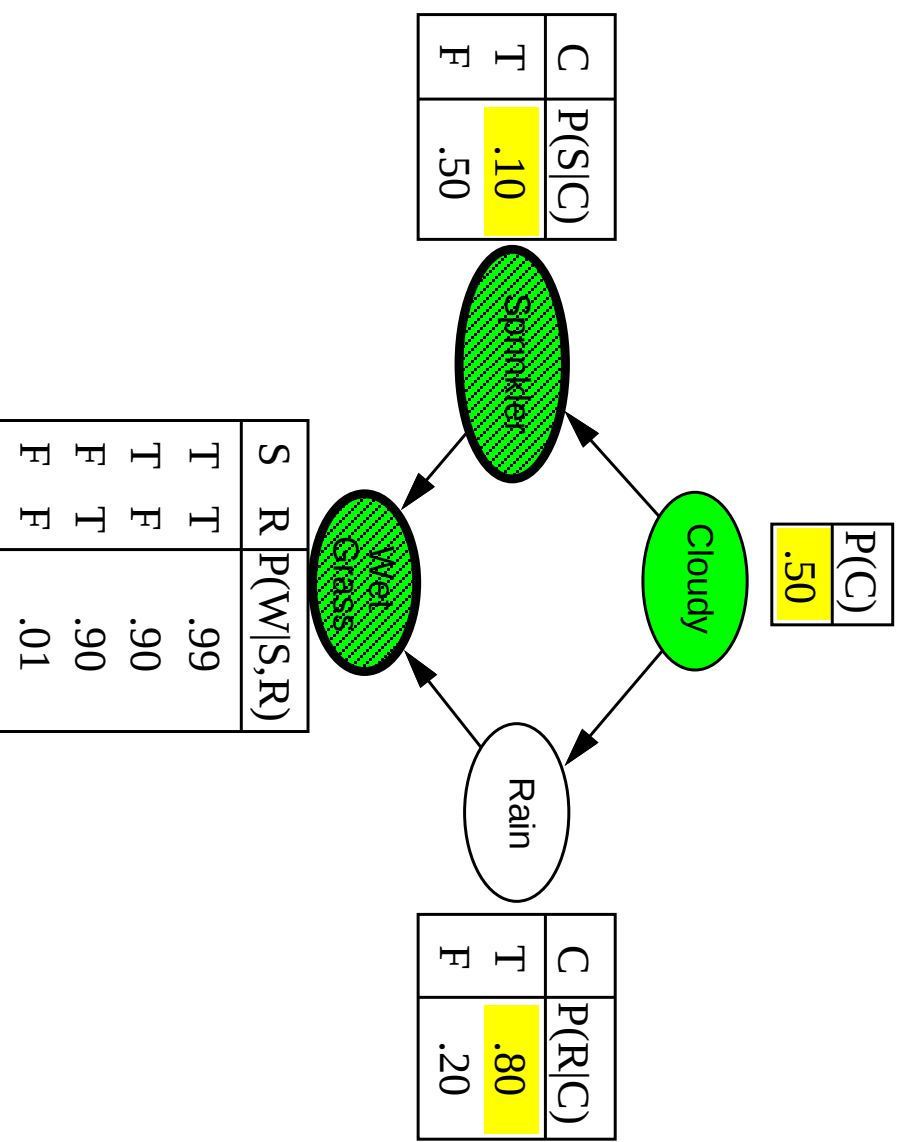
$$w = 1.0$$

Likelihood weighting example



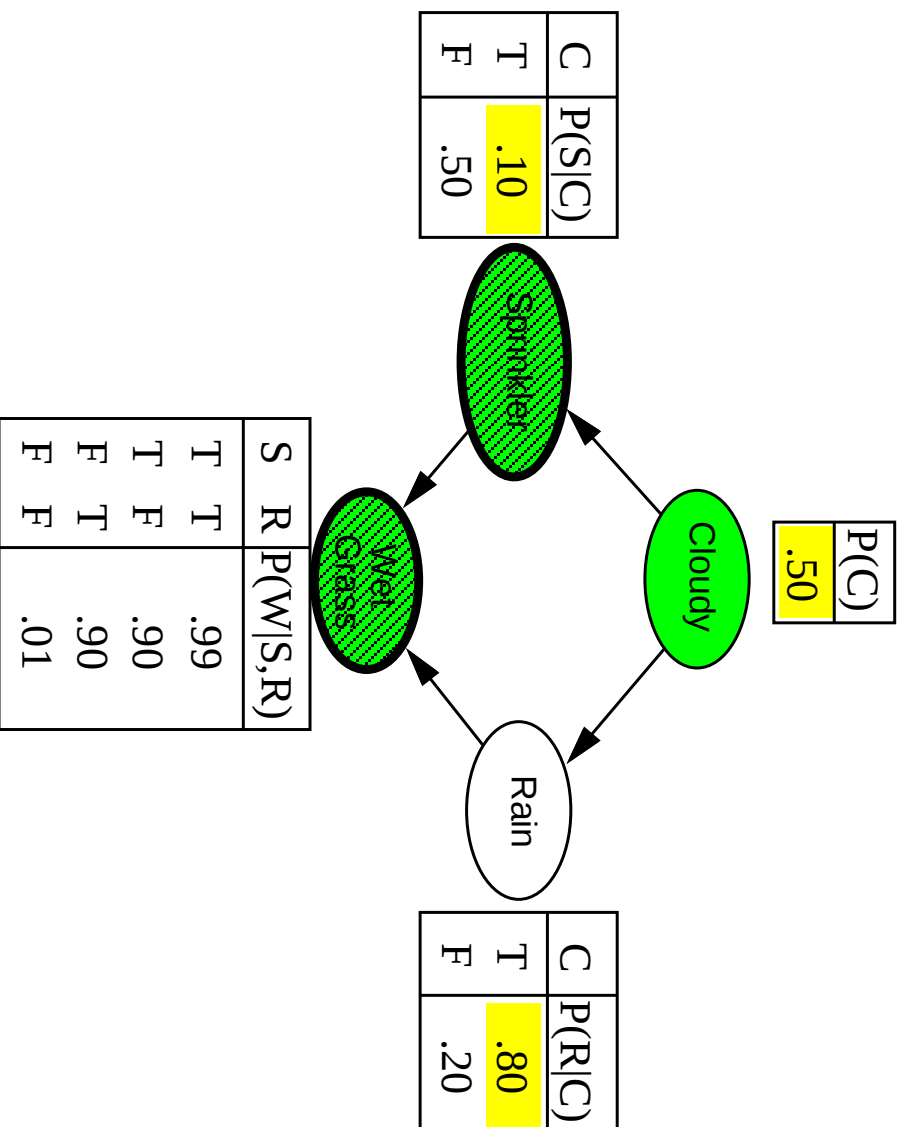
$$w = 1.0$$

Likelihood weighting example



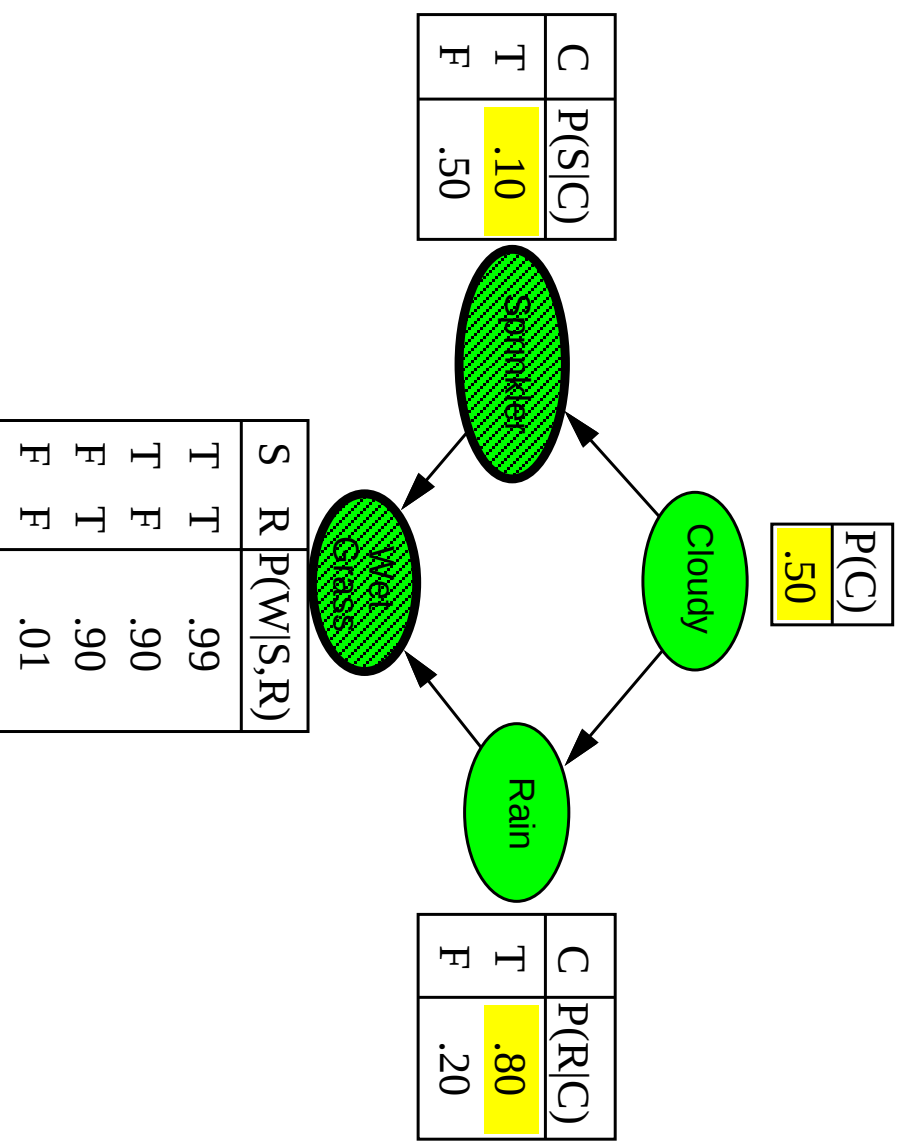
$w = 1.0$

Likelihood weighting example



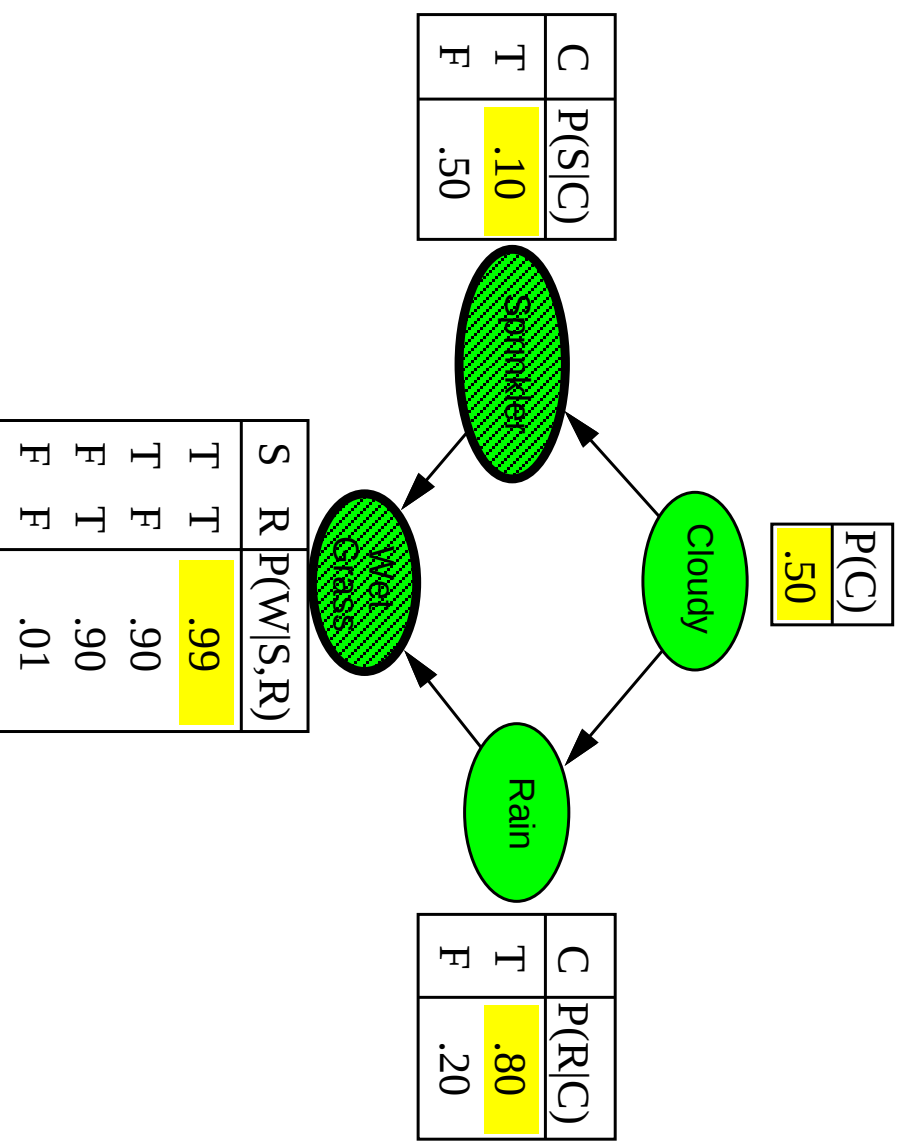
$$w = 1.0 \times 0.1$$

Likelihood weighting example



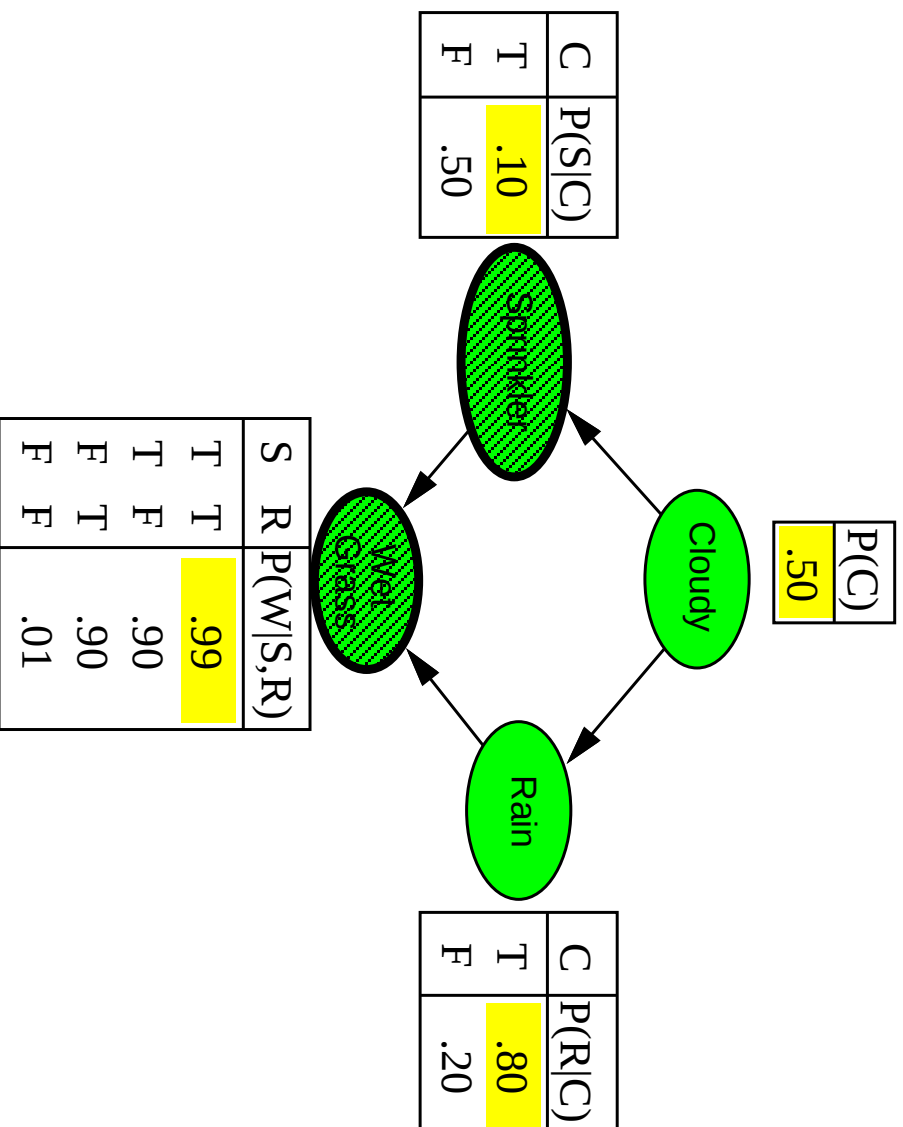
$$w = 1.0 \times 0.1$$

Likelihood weighting example



$$w = 1.0 \times 0.1$$

Likelihood weighting example



$$w = 1.0 \times 0.1 \times 0.99 = 0.099$$

Likelihood weighting analysis

Sampling probability for WEIGHTEDSAMPLE is

$$S_{WS}(\mathbf{z}, \mathbf{e}) = \prod_{i=1}^l P(z_i | Parents(Z_i))$$

Note: pays attention to evidence in **ancestors** only

⇒ somewhere “in between” prior and posterior distribution

Weight for a given sample $\mathbf{z}, \mathbf{e}$ is

$$w(\mathbf{z}, \mathbf{e}) = \prod_{i=1}^m P(e_i | Parents(E_i))$$

Weighted sampling probability is

$$\begin{aligned} S_{WS}(\mathbf{z}, \mathbf{e}) w(\mathbf{z}, \mathbf{e}) \\ &= \prod_{i=1}^l P(z_i | Parents(Z_i)) \prod_{i=1}^m P(e_i | Parents(E_i)) \\ &= P(\mathbf{z}, \mathbf{e}) \text{ (by standard global semantics of network)} \end{aligned}$$

Hence likelihood weighting returns consistent estimates but performance still degrades with many evidence variables because a few samples have nearly all the total weight

